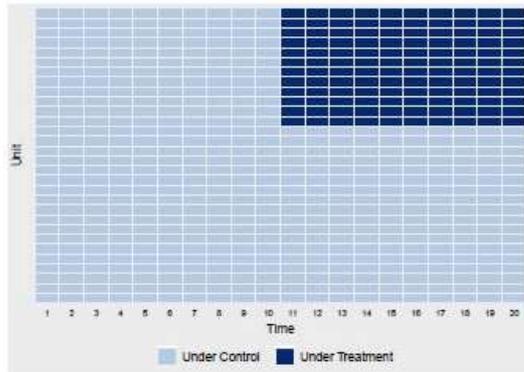
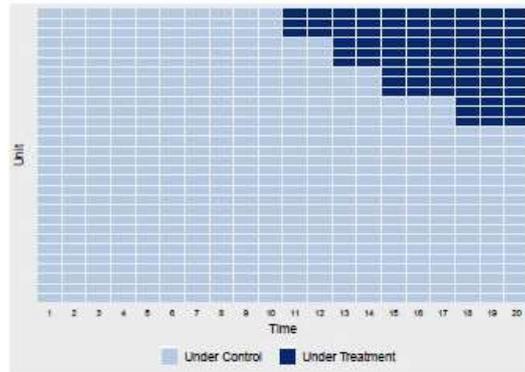


# Time Series Cross Sectional Data Analysis

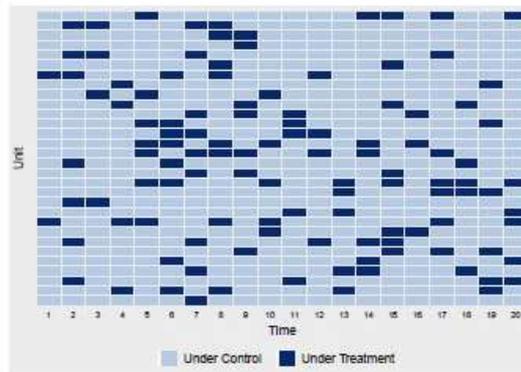
## I. TSCS 형태



(a) Multi-period DID



(b) Staggered Adoption



(c) General Pattern

## II. TSCS 접근방법 1 (강한 외생성: Strict Exogeneity Regime)

<examples>: DID, TWFE, LFM, SCM

1. DM-LFM model 일반적 수식(Pang et al., 2022) = dynamic multilevel latent factor model

$$\begin{aligned}
 y_{it} &= \delta_{it}\omega_{it} + X'_{it}\beta + Z'_{it}\alpha_i + A'_{it}\xi_t + F\gamma' + \epsilon \\
 &= \delta_{it}\omega_{it} + X'_{it}\beta_{it} + \gamma'_i f_t + \epsilon_{it}, \text{ if } \beta_{it} = \beta + \alpha_i + \xi_t, \xi_t = \Phi_\xi \xi_{t-1} + e_t, f_t = \Phi_f f_{t-1} + \nu_t \\
 &= \delta_{it}\omega_{it} + X'_{it}(\beta + \alpha_i + \xi_t) + \gamma'_i f_t + \epsilon_{it} \\
 &= \delta_{it}\omega_{it} + X'_{it}\beta + X'_{it}(\omega_\alpha \cdot \tilde{\alpha}_i) + X'_{it}(\omega_\xi \cdot \tilde{\xi}_t) + (\omega_\gamma \cdot \tilde{\gamma})' f_t + \epsilon_{it}
 \end{aligned}$$

$$\begin{aligned}
 U &= (U_1, U_2, \dots, U_N) = \Gamma' F \\
 \Gamma &= (\gamma_1, \gamma_2, \dots, \gamma_N) (r \times N) (r \ll \min N, T); \Gamma_0 = \text{Diag}() \\
 F &= (f_1, f_2, \dots, f_T)' (r \times T) \\
 X_{it} &: (T \times p_1) \\
 Z_{jt} (j=1, 2, \dots, p_2) &: (T \times p_2) \\
 A_{jt} (j=1, 2, \dots, p_3) &: (T \times p_3) \\
 f_j (j=1, 2, \dots, r) &: (T \times r); \Gamma_0 = \text{Diag}(\omega'_{\gamma_1}, \omega'_{\gamma_2}, \dots, \omega'_{\gamma_r}) \\
 \beta &: (p_1 \times 1) \\
 \alpha_i &: (p_2 \times 1); H_0 = \text{Diag}(\omega'_{\alpha_1}, \omega'_{\alpha_2}, \dots, \omega'_{\alpha_{p_2}}) \\
 \xi &= (\xi'_1, \xi'_2, \xi'_3, \xi'_4, \dots, \xi'_T)' : (p_3 \times 1); \Sigma_e = \text{Diag}(\omega'_{\xi_1}, \omega'_{\xi_2}, \dots, \omega'_{\xi_{p_3}}) \\
 \nu, e, f_t &\sim N(0, 1) \\
 i &= (1, 2, 3, \dots, N) \\
 t &= (1, 2, 3, \dots, T) \\
 \beta_k | \tau_{\beta_k}^2 &\sim N(0, \tau_{\beta_k}^2), \tau_{\beta_k}^2 | \lambda_\beta \sim \text{Exp}\left(\frac{\lambda_\beta^2}{2}\right), \lambda_\beta^2 \sim G(a_1, a_2), k = 1, 2, \dots, p_1
 \end{aligned}$$

2. DiD model(Liu et al., 2020) = fixed effects counterfactual model

$$\begin{aligned}
 Z_i &= A_i = (1, 1, 1, \dots, 1)' \quad \gamma = 0; \\
 \omega_\alpha &= \omega_\beta = \omega_\gamma = 0 \\
 \therefore y_{it} &= \delta_{it}\omega_{it} + X'_{it}\beta + \alpha_i + \xi_t + \epsilon_{it}
 \end{aligned}$$

3. SCM model(Abadie et al., 2010) = a factor model = Synthetic Control Method

$$\begin{aligned}
 Z_{it} &= \emptyset, X_i = A_i \text{ time-invariant}; \\
 \therefore y_{it} &= \delta_{it}\omega_{it} + X'_{it}\beta_t + \xi_t + \gamma'_i f_t + \epsilon_{it} \\
 \implies &\text{비교사례연구에 적합(개체가 적은 특징)}
 \end{aligned}$$

4. Gsyth model(Xu, 2017): Factor-Augmented Approach

$$\begin{aligned}
 Z_{it} &= A_{it} = \emptyset \\
 \therefore y_{it} &= \delta_{it}\omega_{it} + X'_{it}\beta + \gamma'_i f_t + \epsilon_{it}
 \end{aligned}$$

5. TWTE(Angrist & Pischke, 2009)

$$y_{it} = \delta\omega_{it} + X'_{it}\beta + \alpha_i + \xi_t + \epsilon_{it}$$

=>  $\delta$ 가 고정되어 있다는 점에서, DID의  $\delta_{it}$ 와 근본적 차이가 있음  
 $\alpha$ (alpha),  $\beta$ (beta),  $\gamma$ (gamma),  $\delta$ (delta),  $\epsilon$ (epsilon),  $\xi$ (xi, 크시)

6. 참고사항

가. DID

(1) Multi-period DID

- Athey and Imbens(2018)
- Egami & Yamauchi(2021)

(2) Staggered Adoption DID

- Goodman-Bacon (2021)
- Callaway & Sant'Anna (2020)

나. DID Extensions

(1) Semiparametric DID Approach (propensity score model, inverse propensity weight)

- Abadie (2005)
- Strezhnev (2018) : Semiparametric + Staggered Adoption
- Sant'Anna & Zhao (2020): doubly robust DID based on the conditional PTA.

다. Twoway Fixed Effects (TWFE)

- 일반적인 패널데이터(figure C)에 매우 적합(Angrist & Pischke, 2009: 236-243)
- 한계 1:  $\delta$ (처치효과)가 상수라고 가정하면, 강한 외생성(strict exogeneity) 가정을 전제함
- 한계 2: 시간에 따른 역동적인 처치효과를 감안하지 못함. 처치효과가 개체별로 이분산적으로 진화한다면, 처치효과는 편향적이다.
- 한계 3: 지연된 DV의 부재
- 한계 4: no carryover effects.

라. SCM

마. DM-LFM(Factor-Augmented Approach)

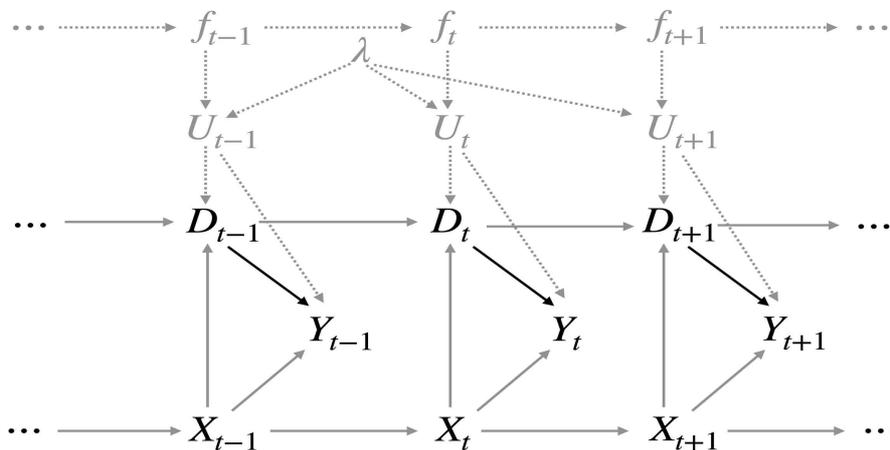


그림 2 DAG for DGPs under Factor-Augmented Approach

## DM-LFM 정리(Strict Exogeneity Regime)

### 1.기본가정. Athey & Imbens(2018)

$$\begin{aligned}
 i &= (1, 2, \dots, N) \\
 t &= (1, 2, \dots, T) \\
 N &= N_{co} + N_{tr} \\
 a_i &\in \{1, 2, \dots, T, c\} \\
 N_{co} \text{ if } a_i = c &> T \\
 N_{tr} \text{ if } (a_i = 1, 2, \dots, T) \\
 \therefore T_{0,i} &= a_i - 1
 \end{aligned}$$

$$\begin{aligned}
 w_i &= (w_{i1}, w_{i2}, \dots, w_{iT})' \\
 w_i : w_i(a_i) &= \begin{cases} w = 0 & \text{if } t < a_i \\ w = 1 & \text{if } t \geq a_i \end{cases} \\
 W_{(N \times T)} &= w_1, w_2, \dots, w_N
 \end{aligned}$$

#### Assumption 1(Cross-Sectional Stable Unit Treatment Value Assumption: SUTVA)

- 횡단면적 파급효과를 배제하고, 잠재적 결과 경로의 수를 크게 줄임
- $$y_{it}(W_{(N \times T)}) = y_{it}(w_i) = y_{it}(w_i(a_i)), \forall i, t$$
- 단위(i)의 잠재적 결과는 단위(i)의 처치상태에 대한 함수로 정의될 수 있음
  - 예컨대, 통독의 효과 연구에서, 통일이 다른 국가의 경제성장에 영향을 미친다는 가정을 배제하는 것임(이것은 매우 강한 가정임)
  - 다른 한편, 선거제도 변경 효과 연구에서, A 주(state)의 선거제도 변경 법률의 채택은 B 주의 선거제도 변경 법률안의 채택여부와 상관없이 B주의 투표율에 영향을 미치지 않는다.

#### [Assumption 1 + Assumption 4 => Strict Exogeneity 엄격한 외생성]

$$\{ Y_{it}(0), Y_{it}(1) \} \perp\!\!\!\perp D_{is} \mid X_i^{1:T}, \alpha_i, f^{1:T}, \forall i, t, s$$

$D_{is}$ : 단위(i)가 시간(s)에서의 처치상태(통제집단 또는 처치집단 여부)

$Y_{it}(0)$ : 통제집단 상태에서의 잠재적 결과  $Y_{it}(1)$ : 처치집단 상태에서의 잠재적 결과

$\therefore$  처치집단 선정은 기준선에서 이미 결정되며, 결과 실현과는 무관하다.

- 과거 결과가 현 결과에 영향 없음, 피드백 효과없음, 이월효과 없음 가정함

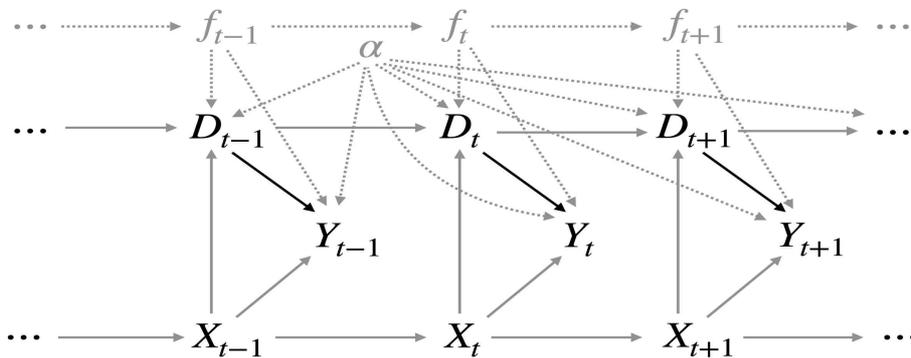


그림 3 DAG for DGPs Under Strict Exogeneity

### Assumption 2(No Anticipation)

$$y_{it}(a_i) = y_{it}(c), \text{ for } t < a_i, \forall i$$

counterfactual outcome:  $y_{it}(c)$

Estimands

$$\delta_{it} = y_{it}(a_i) - y_{it}(c), \text{ for } a_i \leq t \leq T$$

ATT(average treatment effect on the treated) for a duration of p periods:

$$\delta_p = \frac{1}{N_{tr,p}} \sum_{i: T-p+1 \leq a_i \leq T} \delta_{i, a_i+p-1}$$

$Y(0)_{(N \times T)}$  under  $W = 0$  (i.e.,  $a_i = c, \forall i$ )

$$Y(0)_{(N \times T)} \in \begin{cases} S_0 \equiv \{(it) | w_{it} = 0\} y_{it}(c) \text{ observed} \\ S_1 \equiv \{(it) | w_{it} = 1\} y_{it}(c) \text{ missing} \end{cases}$$

$$\therefore Y(0) = Y(0)^{obs} + Y(0)^{mis}$$

$$X_{(T \times p_1)} = \{X_1, X_2, \dots, X_N\}, X_i = (X_{i1}, X_{i2}, \dots, X_{iT})'$$

### 2. Assignment Mechanism:Rubin et al.(2010)

$$Pr(Y(W)^{mis} | X, Y(W)^{obs}, W)$$

$$\begin{aligned} Pr(Y(0)^{mis} | X, Y(0)^{obs}, A) &= \frac{Pr(X, Y(0)^{mis}, Y(0)^{obs}) Pr(A | X, Y(0)^{mis}, Y(0)^{obs})}{Pr(X, Y(0)^{obs}, A)} \\ &\propto Pr(X, Y(0)^{mis}, Y(0)^{obs}) Pr(A | X, Y(0)^{mis}, Y(0)^{obs}) \\ &\propto Pr(X, Y(0)) Pr(A | X, Y(0)) \end{aligned}$$

Assumption 3(individualistics assignment and positivity) no policy diffusion effects

$$Pr(A | X, Y(0)) = \prod_{i=1}^n Pr(a_i | X_i, Y_i(0)), 0 < Pr(a_i | X_i, Y_i(0)) < 1 \forall i$$

Assumption 4(Latent ignorability) implies parallel trends assumptions

$$\text{a vector of latent variables } U_i = (u_{i1}, u_{i2}, \dots, u_{iT}), u_{it} = \gamma_i \cdot g(t)$$

$$Pr(a_i | X_i, Y_i(0), U_i) = Pr(a_i | X_i, Y_i(0)^{mis}, Y_i(0)^{obs}, U_i) = Pr(a_i | X_i, U_i)$$

==> we will extract  $U_{it}$  from  $Y_i(0)^{obs}$ .

parallel trends assumptions:  $u_{i1} = u_{i2} = u_{i3} = \dots = u_{iT} = u_i$

Assumption 5 (Feasible data extraction)

$$U_{(N \times T)} = (U_1, U_2, \dots, U_N)$$

$$U = \Gamma F$$

$$F_{(r \times T)} = (f_1, f_2, \dots, f_T)$$

$$\Gamma_{(r \times N)} = (\gamma_1, \gamma_2, \dots, \gamma_N)$$

### 3. Posterior Predictive Inference

$X'$ ;  $X$  &  $U$

$$\begin{aligned} Pr(Y(0)^{mis} | X', Y(0)^{obs}, A) &= \frac{Pr(X', Y(0)^{mis}, Y(0)^{obs})Pr(A | X', Y(0)^{mis}, Y(0)^{obs})}{Pr(X', Y(0)^{obs}, A)} \\ &\propto Pr(X', Y(0)^{mis}, Y(0)^{obs})Pr(A | X') \\ &\propto Pr(X', Y(0)) \end{aligned}$$

Assumption 6 (Exchangeability)

$\{(X'_{it}, y_{it}(c))\}$  is invariant to permutations in the index "it".

By de Finetti's theorem(de Finetti, 1963)

$$\begin{aligned} Pr(Y(0)^{mis} | X', Y(0)^{obs}, A) &\propto Pr(\{(X'_{it}, y_{it}(c))\}) \\ &\propto \int \left( \prod_{it \in S_1} f(y_{it}(c)^{mis} | X_{it}, \theta') \right) \left( \prod_{it \in S_0} f(y_{it}(c)^{obs} | X_{it}, \theta') \right) \pi(\theta) d\theta \\ &\text{<posterior predictive distribution> <likelihood>} \end{aligned}$$

$\theta$ : the parameters that govern the DGP of  $y_{it}(c)$  given  $X'_{it}$ ,  $\theta' = (\theta, U)$

### 4. Estimation Strategy(추정방법)

$O = \{(i, t) | D_{it} = 0\}$ ,  $M = \{(i, t) | i \in T, D_{it} = 1\}$  O=Observed, M=Missing

- (1) "O"(통제집단)를 이용하여,  $Y_{it}(0) = f(X_{it}) + h(U_{it}) + \epsilon_{it}$ 를 추정하는  $\hat{f}$ ,  $\hat{h}$ 를 구함
- (2) 처치집단 개체에 대해 반사실적 결과( $Y_{it}(0)$ )를 예측함

$$\hat{Y}_{it}(0) = \hat{f}(X_{it}) + \hat{h}(U_{it}) \text{ for all } (i, t) \in M$$

- (3)  $\delta_{it}$ 를 추정함  $\hat{\delta}_{it} = Y_{it} - \hat{Y}_{it}(0)$  for all treated observation  $(i, t) \in M$
- (4)  $\hat{\delta}_{it}$ 를 활용하여, ATT, ATT(s)를 추정함

$$\widehat{ATT} = \frac{1}{|M|} \sum_M \hat{\delta}_{it},$$

$$\widehat{ATT}_s = \frac{1}{|S|} \sum_{(i,t) \in S} \hat{\delta}_{it}, S = \{(i, t) | D_{i,t-s} = 0, D_{i,t-s+1} = D_{i,t-s+2} = \dots = D_{it} = 1\},$$

### III. TSCS 접근방법 2 (순차적 무시성: Sequential Ignorability Regime)

#### 1. Key Assumption

$$\{Y_{it}(0), Y_{it}(1)\} \perp\!\!\!\perp D_{it} \mid X_i^{1:t}, Y_i^{1:t-1}, \forall i, t$$

$X_i^{1:t}$ : 시간(t)까지의 공변인들의 역사

$Y_i^{1:(t-1)} = \{Y_{i1}, Y_{i2}, \dots, Y_{i,(t-1)}\}$ : (t-1) 시간까지의 종속변수의 역사

- 순차적 무시성; 처치집단 선정은 모든 과거 정보(공변인과 종속변수 포함)를 포함하여 무시할 수 있음
- 피드백 효과 인정( $Y_{t-1} \Rightarrow D_t, X_{t-1} \Rightarrow D_t$ )
- 과거의 결과가 직접적으로 현재의 결과에 영향을 미침( $Y_{t-1} \Rightarrow Y_t$ )
- 이월효과 인정( $D_{t-1} \Rightarrow Y_t$ )

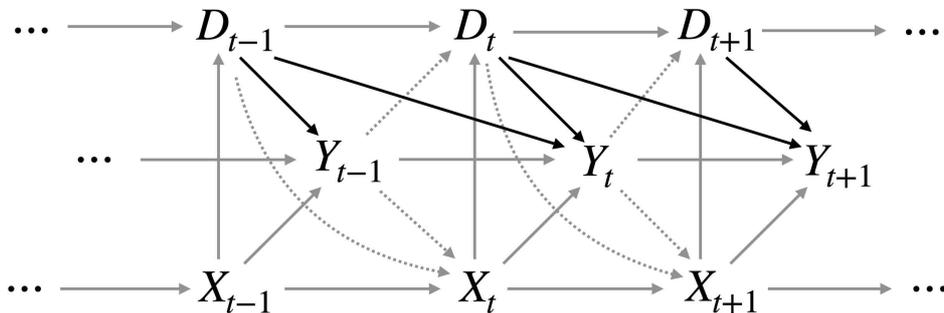


그림 4 DAG for DGPs under Sequential Ignorability

#### 2. Ideal experiment: sequential randomization

3. examples: Lagged Dependent Variable Models(LDV), Autoregressive Distributed Lag Models(ADL), Marginal Structural Models(MSM)

#### 4. PanelMatch (Imai, Kim & Wang, 2021)

- 1 step: find matched units
- 2 step: matching or reweighting
- 3 step: block bootstrapping. compute ATT
- 한계는, 데이터가 줄어들음

#### 4. MSM(Marginal Structural Models) - epidemiology & biomedical sciences

- Blackwell & Glynn(2018), Kurer (2020)